

Data-driven Prediction and Control Through the Lens of System Identification

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Trends on Dissipativity in Systems and Control - Brig - May 23, 2022



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Paradigm shift in Engineering

- From component-oriented to system-level design (smart cities, intelligent transportation, industrial automation..)
- From first-principles modelling to data-driven

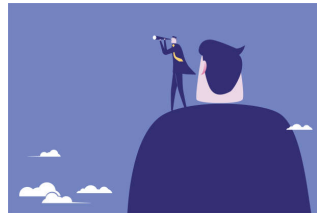


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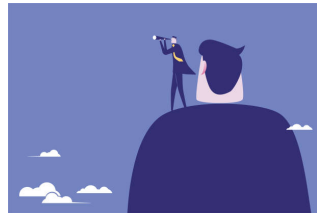
- From component-oriented to system-level design (smart cities, intelligent transportation, industrial automation..)
- From first-principles modelling to data-driven
- **Methods** and **tools** must adapt, but not the **foundational principles**



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(optimal, robust, adaptive control,
system identification, dynamical
systems theory)..



- Standing on the shoulders of **giants**
(optimal, robust, adaptive control,
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- ..seeking new peaks (statistical learning,
optimization, information theory)



(New) Direct data-driven methods through the lens of a(n old) giant

Methods for simulating and controlling dynamical systems **directly** from data

→ no knowledge of an explicit model is required

- Behavioral system theory (noise-free data)

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- Recent works towards a stochastic counterpart [1-2]

¹ Pan Ou Faulwasser, *"On a Stochastic Fundamental Lemma and Its Use for Data-Driven MPC"* 2021

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Joint work with

- Behavioral system theory (noise-free data)
- Recent works towards a stochastic counterpart [1-2]
- Today: System identification-based statistical reasoning
- Framework for simulation, filtering and control from noisy data



Mingzhou Yin



Roy S. Smith

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- 1 A Bayesian perspective on the Fundamental Lemma
- 2 Confidence regions for noisy predictions
- 3 Informative data for simulation problems

Representation-free perspective on dynamical systems as sets of trajectories [Willems, 1987]

- The *Fundamental Lemma* [Willems, 2005] arguably a driving force behind recently proposed subspace-type data-driven analysis and control methods
- Non-parametric representation of an LTI system by the image of a Hankel data matrix

$$z_{[i,j]} \rightarrow \mathcal{H}_L(z_{[i,j]}) := \begin{bmatrix} z_i & z_{i+1} & \cdots & z_{N+i-L} \\ z_{i+1} & z_{i+2} & \cdots & z_{N+i-L+1} \\ \vdots & \vdots & & \vdots \\ z_{i+L-1} & z_{i+L} & \cdots & z_{N+i-1} \end{bmatrix}$$

$z_{[i,j]}$ is: **persistently exciting** (PE) of order L if $\mathcal{H}_L(z_{[i,j]})$ has full row rank

- Can we use raw data to span the I/O behavior of the system?

Spanning the behavior?

Minimal linear time-invariant system

$$x_{t+1} = Ax_t + Bu_t$$

$$y_t = Cx_t + Du_t$$

$x \in \mathbb{R}^{n_x}$ is the state, $u \in \mathbb{R}^{n_u}$ is the input, and $y \in \mathbb{R}^{n_y}$ is the output (the *lag* l is the smallest integer i such that the ext. observability matrix $\mathcal{O}_i$ has rank n_x).

Direct prediction/simulation problem

Given: an input-output **data trajectory** $(u_d [0, N-1], y_d [0, N-1])$.

Goal: for any *input simulation trajectory* $u_s [0, L_s-1]$, and any initial condition, find the (unique) *output simulation trajectory* $y_s [0, L_s-1]$.

[Willems, 2005]

Controllable LTI system with McMillan degree n_x , I/O trajectory $(u_d [0, N-1], y_d [0, N-1])$. If $u_d [0, N-1]$ is PE of order $L + n_x$, then $(u_t [0, L-1], y_t [0, L-1])$ is an I/O trajectory **iff** there exists g s.t.

$$\begin{bmatrix} u_t [0, L-1] \\ y_t [0, L-1] \end{bmatrix} = \begin{bmatrix} \mathcal{H}_L(u_d [0, N-1]) \\ \mathcal{H}_L(y_d [0, N-1]) \end{bmatrix} g$$

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Given $L_0 \geq l$ (define $L_s = L - L_0$), $y_s [0, L_s-1]$ (y_s) is the unique output trajectory with past trajectory $(u_{ini} [0, L_0-1], y_{ini} [0, L_0-1])$ (u_{ini}, y_{ini}) and input trajectory $u_s [0, L_s-1]$ (u_s), **iff** there exists g s.t.

$$\begin{bmatrix} u_{ini} \\ u_s \\ y_{ini} \\ y_s \end{bmatrix} = \begin{bmatrix} U \\ Y \end{bmatrix} g = \begin{bmatrix} U_p \\ U_f \\ Y_p \\ Y_f \end{bmatrix} g = \begin{bmatrix} \mathcal{H}_L(u_d [0, N-1]) \\ \mathcal{H}_L(y_d [0, N-1]) \end{bmatrix} g$$

Noise-free case abstraction

The trajectory can be partitioned into a known and unknown part

$$\begin{bmatrix} \hat{z} \\ z \end{bmatrix} = \begin{bmatrix} u \\ y \end{bmatrix}, \quad Z = \begin{bmatrix} U \\ Y \end{bmatrix}, \quad \left(\text{e.g. } \hat{z} = \begin{bmatrix} u_{ini} \\ u_s \\ y_{ini} \end{bmatrix}, z = \begin{bmatrix} y_s \end{bmatrix} \right)$$

$$\hat{z} = Z_1 g \rightarrow g^*(\hat{z}, Z_1), \quad \text{from: } \underline{\text{known}} \text{ and } data, \text{ to } g$$

$$z = Z_2 g^* = Z_2 g^*(\hat{z}, Z_1) \quad \text{from: } g \text{ and } data, \text{ to } \underline{\text{unknown}}$$

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$$z = \mathcal{F}_Z(\hat{z}), \quad (y_s = \mathcal{F}_Z(u_{ini}, u_s, y_{ini}) \text{ input-output mapping})$$

Noisy case *regression* viewpoint

$$\tilde{y}_i^d = y_i^d + w_i^d, \quad w_i^d \sim \mathcal{N}(0, \sigma_y^2 I_p)$$

$$\tilde{u}_i^d = u_i^d + v_i^d, \quad v_i^d \sim \mathcal{N}(0, \sigma_u^2 I_m)$$

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$$z \sim \mathcal{N}(Zg, \Sigma_z(g)) \quad \text{"Fundamental Lemma"-prior}$$

$$w \sim \mathcal{N}(0, \Sigma_w) \quad \text{Disturbance model}$$

$$\hat{z}|z \sim \mathcal{N}(z, \Sigma_w) \quad \text{Likelihood}$$

$$\Sigma_z(g) = \begin{bmatrix} \Sigma_{U_p}(g) & 0 & 0 & 0 \\ 0 & \Sigma_{U_f}(g) & 0 & 0 \\ 0 & 0 & \Sigma_{Y_p}(g) & 0 \\ 0 & 0 & 0 & \Sigma_{Y_f}(g) \end{bmatrix}$$

The Signal Matrix Model (SMM)

The maximum-a-posteriori (MAP) estimate of the trajectory is

$$\begin{aligned} z^*(g) &= \arg \max_z \underbrace{p(z|\hat{z})}_{\propto p(\hat{z}|z)p(z)} \\ &= \Sigma_z(g) (\Sigma_z(g) + \Sigma_w)^{-1} \hat{z} + \Sigma_w (\Sigma_z(g) + \Sigma_w)^{-1} Zg \end{aligned} \quad (\text{MAP})$$

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where g is obtained through maximum marginal-likelihood (ML)

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ML: from known and *data*, to $g \rightarrow$ MAP: from g and *data*, to unknown.

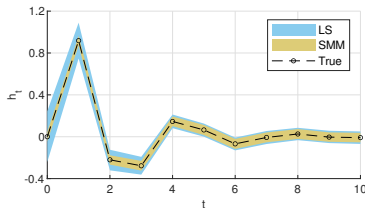
Example 1: Direct data-driven simulation

$$\hat{z} = \begin{bmatrix} u_{ini} \\ u_s \\ y_{ini} \\ 0 \end{bmatrix}, \quad z = \begin{bmatrix} 0 \\ 0 \\ 0 \\ y_s \end{bmatrix}, \quad \Sigma_w = \begin{bmatrix} \Sigma_{u_i} & 0 & 0 & 0 \\ 0 & \Sigma_{u_s} & 0 & 0 \\ 0 & 0 & \Sigma_{y_i} & 0 \\ 0 & 0 & 0 & \Sigma_{y_s} \end{bmatrix}, \quad \Sigma_{u_i} = \Sigma_{u_s} = \Sigma_{y_i} = 0, \quad \Sigma_{y_s} = \infty \rightarrow y_s = Y_f g^*$$

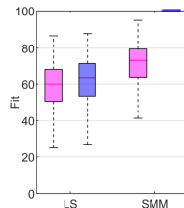
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Truncated Infinite Impulse Response (IIR) estimation, comparison with Least-Squares



Estimates within two standard deviations



Magenta: noisy data; blue: noise-free data

Example 2: Direct data-driven control

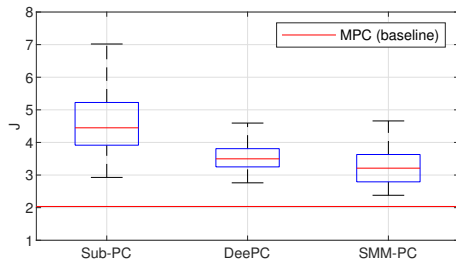
$$\begin{aligned} \min_{\hat{u}, \hat{y}} \quad & \sum_{k=0}^{L_s-1} \left(\left\| \hat{y}_k - y_k^{\text{ref}} \right\|_Q^2 + \left\| \hat{u}_k - u_k^{\text{ref}} \right\|_R^2 \right) \\ \text{s.t.} \quad & \left(\begin{bmatrix} u_{\text{ini}} \\ \hat{u} \end{bmatrix}, \begin{bmatrix} y_{\text{ini}} \\ \hat{y} \end{bmatrix} \right) \text{ is a trajectory,} \\ & \hat{u} \in \mathcal{U}, \hat{y} \in \mathcal{Y}, \end{aligned}$$

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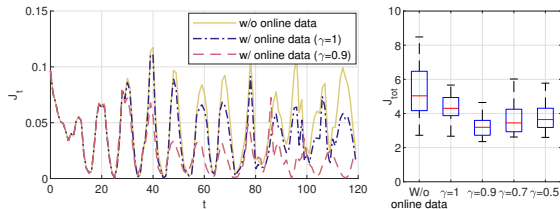
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CL cost comparison with *competitors*



Adaptation using online data

What is the error associated with predicting y_s from noisy data?

Confidence regions for data-driven prediction

What is the error associated with predicting y_s from noisy data?

Prototype of stochastic data-driven predictors

Existing algorithms ^a can be expressed as

$$\begin{aligned} \mathcal{F}_Z(u_{\text{ini}}, u_s, y_{\text{ini}}) = Y_f \operatorname{argmin}_{g(\delta)} \quad & \|\delta\|_2^2 + \lambda \|g\|_2^2 \\ \text{s.t.} \quad & \begin{bmatrix} U_p \\ U_f \\ Y_p \end{bmatrix} g = \begin{bmatrix} u_{\text{ini}} \\ u_s \\ y_{\text{ini}} + \delta \end{bmatrix} \end{aligned} \quad (\text{PRED})$$

where λ is algorithm-dependent.

^aSubspace predictor [Fiedler, 2021], SMM, Wasserstein distance minimization [Lian, 2021]

Goal: obtain **probabilistic** and **deterministic** confidence regions

SMM as data-driven predictor

- Output noise $(y_{\text{ini}}, Y_p, Y_f)$
- Specific matrix structure: M length- L input-output trajectories

$$z_i^d = \text{col} (u_{t_i}^d, \dots, u_{t_i+L-1}^d, y_{t_i}^d, \dots, y_{t_i+L-1}^d) \in \mathbb{R}^{L(n_u+n_y)}, \quad i = 0, \dots, M-1$$

Matrix $Z = \begin{bmatrix} z_0^d & \dots & z_{M-1}^d \end{bmatrix}$ is mosaic Hankel ($t_{i+1} = t_i + 1$) or Page ($t_{i+1} = t_i + L$)

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The ML estimate of g solves

$$\min_g \log \det \left(\begin{bmatrix} \Sigma_{Y_p}(g) & 0 \\ 0 & \Sigma_{Y_f}(g) \end{bmatrix} + \begin{bmatrix} \Sigma_{y_i} & 0 \\ 0 & 0 \end{bmatrix} \right) + \delta^\top \left(\begin{bmatrix} \Sigma_{Y_p}(g) & 0 \\ 0 & \Sigma_{Y_f}(g) \end{bmatrix} + \Sigma_{y_i} \right)^{-1} \underbrace{\delta}_{Y_p g - y_{ini}},$$

where $\Sigma_{y_i} = \sigma_{p,y}^2 I_{n_y L_0}$, $\Sigma_{Y_p}(g) = \sigma_{p,y}^2 \|g\|_2^2 I_{n_y L_0}$, $\Sigma_{Y_f}(g) = \sigma_y^2 \|g\|_2^2 I_{n_y L_s}$.

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Quadratic approximation of the obj. about $g = \hat{g}$ gives (PRED) with $\lambda = n_y \left(L_0 \sigma_{p,y}^2 / \|\hat{g}\|_2^2 + L \sigma_y^2 \right)$.

Probabilistic and robust (in probability) bounds for PRED

Consider prediction $y = \mathcal{F}_Z(\cdot)$ of true output y_0 . Then

$$y - y_0 \mid g, \delta \sim \mathcal{N} \left(\underbrace{\Gamma \delta, \begin{bmatrix} -\Gamma & I_{n_y L'} \end{bmatrix} \begin{bmatrix} \Sigma_{Y_p}(g) & 0 \\ 0 & \Sigma_{Y_f}(g) \end{bmatrix} \begin{bmatrix} -\Gamma^\top \\ I_{n_y L'} \end{bmatrix}}_{\Sigma} + \Gamma \Sigma_{y_i} \Gamma^\top \right)$$

where $\Gamma = \text{col} (CA^{L_0}, \dots, CA^{L-1}) \text{col} (C, \dots, CA^{L_0-1})^\dagger$.

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The true output y_0 belongs w.p. p to

$$\mathcal{Y} = \left\{ \tilde{y} \mid (y - \tilde{y} - \Gamma \delta)^\top \Sigma^{-1} (y - \tilde{y} - \Gamma \delta) \leq \mu_p \right\}, \quad F_{\chi^2(L')}(\mu_p) = p$$

$F_{\chi^2(d)}(\cdot)$ is the cumulative distribution function of the χ^2 -distribution with d degrees of freedom.

Insights from the proof

- $y - y_0 = \underbrace{E_f g}_{\text{noise in } Y_f} + \underbrace{y^-}_{\text{wrong initial condition}}$

$$E_f = Y_f - Y_f^0; \quad y^- \text{ free response from I.C. } u_{\text{ini}}^- = 0, y_{\text{ini}}^- = Y_p^0 g - y_{\text{ini}}^0$$

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- $y_{\text{ini}}^- = \delta + (y_{\text{ini}} - y_{\text{ini}}^0) - E_p g$

$$y^- = \begin{bmatrix} CA^{L_0} \\ \vdots \\ CA^{L-1} \end{bmatrix} x^-, \quad y_{\text{ini}}^- = \begin{bmatrix} C \\ \vdots \\ CA^{L_0-1} \end{bmatrix} x^- \Rightarrow y^- = \Gamma y_{\text{ini}}^-$$

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 $E_f = Y_f - Y_f^0$; y^- free response from I.C. $u_{\text{ini}}^- = 0$, $y_{\text{ini}}^- = Y_p^0 g - y_{\text{ini}}^0$
- $y_{\text{ini}}^- = \delta + (y_{\text{ini}} - y_{\text{ini}}^0) - E_p g$

$$y^- = \begin{bmatrix} CA^{L_0} \\ \vdots \\ CA^{L-1} \end{bmatrix} x^-, \quad y_{\text{ini}}^- = \begin{bmatrix} C \\ \vdots \\ CA^{L_0-1} \end{bmatrix} x^- \Rightarrow y^- = \Gamma y_{\text{ini}}^-$$

- Holds for general predictors

$$\mathcal{F}_Z(\cdot) = Y_f g, \quad \text{s.t.} \quad \begin{bmatrix} U_p \\ U_f \\ Y_p \end{bmatrix} g = \begin{bmatrix} u_{\text{ini}} \\ u_s \\ y_{\text{ini}} + \delta \end{bmatrix}$$

- ✗ Confidence regions depend on **model's quantities** (range space of observability);
no closed form solution in more general (noise and matrix structures) cases

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no closed form solution in more general (noise and matrix structures) cases
- ✓ Quantification of the prediction error due to noisy information;
tool for designing **new predictors**

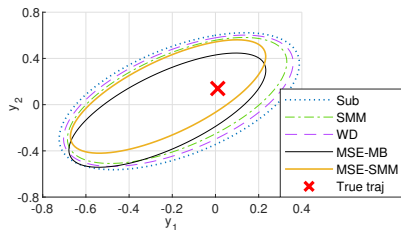
Minimum mean-squared error predictor

The minimum MSE predictor of the class (PRED) is

$$\mathcal{F}_Z(\cdot) = Y_f \operatorname{argmin}_{g(\cdot, \delta)} \underbrace{\delta^\top \Gamma^\top \Gamma \delta + \operatorname{tr}(\Sigma)}_{\text{MSE}(g, \delta)} = \operatorname{argmin}_{g(\cdot, \delta)} \|\delta\|_Q^2 + \lambda_{\text{MSE}} \|g\|_2^2$$
$$\text{s.t.} \quad \begin{bmatrix} U_p \\ U_f \\ Y_p \end{bmatrix} g = \begin{bmatrix} u_{\text{ini}} \\ u_s \\ y_{\text{ini}} + \delta \end{bmatrix} \quad (\text{MSE-PRED})$$

Numerical experiments

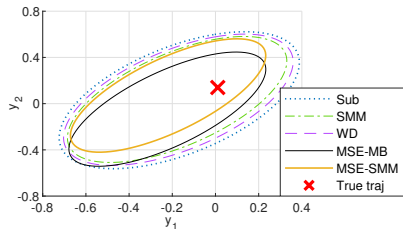
Two-output 4th order system. $L = 10$, $L_0 = 8$,
 $M = 80$, $\sigma^2 = 0.1$, $\rho = 0.90$, unit Gaussian input.



CR of different predictors (based on true Γ).

Numerical experiments

Two-output 4th order system. $L = 10$, $L_0 = 8$,
 $M = 80$, $\sigma^2 = 0.1$, $\rho = 0.90$, unit Gaussian input.



CR of different predictors (based on true Γ).

1000 randomly generated systems (states between 3 and 8). $L = 20$, $L_0 = 8$, $L' = 12$,
 $M = 320$, unit Gaussian input.

Comparison of the empirical MSE.

	$\sigma^2 = 0.1$	$\sigma^2 = 0.5$	$\sigma^2 = 1$
Sub	0.115	0.558	1.106
SMM	0.099	0.476	0.915
WD	0.113	0.548	1.091
MSE-SMM	0.096	0.460	0.897
MSE-MB	0.094	0.435	0.833

How to best choose u_d ?

- The *Fundamental Lemma* is effectively an **experiment design** result
- Goal: go beyond PRBS (Pseudo Random Binary Signals) and randomly selected input..
- **Data informativity** for the *noisy* given-input simulation problem
- Informativity encompasses both the selection of the input and the matrix structure (Hankel/Page)

Bayesian design of experiment (DOE)

Interpret the output trajectory y_s as a random variable.

$$y_s^1 \sim \mathcal{N}(0, \Sigma_K) \quad \rightarrow \quad \textit{Prior distribution}$$

$$y_s^2 \sim \mathcal{N}(Y_f g^*, \underbrace{\Sigma_{Y_f}(g^*)}_{\text{sub-matrix of } \Sigma_z(g^*)}) \quad \rightarrow \quad \textit{Data-based estimate (via SMM)}$$

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Combine prior and data-based estimates to provide the posterior

$$(y_s | (u_d, y_d)) \sim \mathcal{N}(K(Y_f g^*), \Sigma_{\text{post}}) \rightarrow \text{Posterior of } y_s$$

where $K = \Sigma_K (\Sigma_K + \Sigma_{Y_f})^{-1}$ is the Kalman gain and $\Sigma_{\text{post}} = \Sigma_K - \Sigma_K (\Sigma_K + \Sigma_{Y_f})^{-1} \Sigma_K$ is the posterior covariance.

The experiment is **meaningful** when the *difference* between prior and posterior distributions is **maximized**.

Mutual information [Cover, 1991] of two multivariate random variables x and y

$$I(x; y) = H(x) + H(y) - H(x, y) = H(x) - H(x|y)$$

$$\Rightarrow I(y_s; (u_d, y_d)) = \mathbb{E} \left[\underbrace{D(y_s | (u_d, y_d) || p(y_s))}_{\text{KL divergence between prior and posterior trajectory}} \right]$$

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Information criterion for SMM

If $\Sigma_{Y_f}(g) = \sigma_y^2 \|g\|_2^2 I_{n_y L_s}$, then there exists a function f such that

$$I(y_s; (u_d, y_d)) = f(\|g\|_2^2, \Sigma_K)$$

where $f(\|g\|_2^2, \cdot)$ is monotonically decreasing.

The input data trajectory can be designed to maximize the information:

$$u_d [0, N-1] \rightarrow U, Y \rightarrow g^* \rightarrow I$$

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$$\min_{g, u_d} \|g\|_2^2 \rightarrow \text{information criterion}$$

$$\text{s.t. } u_d \in \mathcal{U} \rightarrow \text{input constraints}$$

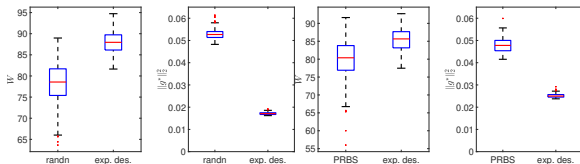
$$g \in \arg \min_{g \in \mathcal{G}, u_d} \log \det(\Sigma_y(g)) + \begin{bmatrix} Y_p g - y_{ini} \\ 0 \end{bmatrix}^\top \Sigma_y^{-1}(g) \begin{bmatrix} Y_p g - y_{ini} \\ 0 \end{bmatrix} \rightarrow \text{ML estimator}$$

Two **challenges** associated with the bi-level program:

- Inner-problem depends on Y_p (implicitly function of u_d)
- Dealing with the bi-level program

Truncated IIR estimation with Hankel matrices

Fit metric: $W = 100 \left(1 - \left[\frac{\sum_{i=0}^{L_s-1} (y_{s,i} - \hat{y}_{s,i})^2}{\sum_{i=0}^{L_s-1} (y_{s,i} - \bar{y}_s)^2} \right]^{1/2} \right)$



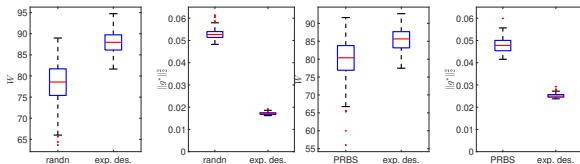
Energy constraint.

Magnitude constraint.

- Optimally designed input perform better than standard PE inputs
- Informativity key to achieve the improvement

Truncated IIR estimation with Hankel matrices

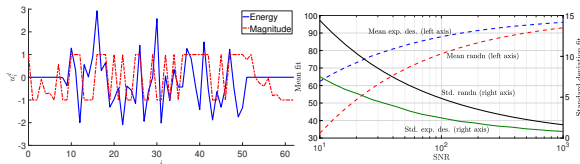
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Energy constraint.

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Optimal inputs.

Fit as a function of SNR.

- Optimal input do not have “intuitive” shape
- Impact of DOE is a function of the SNR

- Statistical framework for the direct prediction and control problems
- First prediction error bounds and data informativity studies in this implicit non-parametric setting
- Towards a principled way to reason about robustness in data-driven problems

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- Stochastic MPC formulation based on prediction error-bounds
- Informative input for control (dual control in direct problems)**
- Adaptive signal matrices: when and why?

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Thank you for your attention!